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Foreign direct investment, diversification and firm performance

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less harmful for diversified (multi-segment) than specialized (single-segment) firms. The larger gains to diversified firms suggest that operational and internal capital market efficiency gains are considerably greater in multi-segment than single-segment firms when both expand their core business overseas.

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- 4. Even though foreign direct investments are subject to many of the same influences and motivations as domestic investments (e.g. <u>Agmon and Lessard</u>, 1977; <u>Errunza and Senbet</u>, 1981; <u>Adler and Dumas</u>, 1983; <u>Fatemi</u>, 1984; <u>Hisey and Caves</u>, 1985; <u>Doukas and Travlos</u>, 1988; <u>Morck and Yeung</u>, 1991, 1992; <u>Williamson</u>, 1970), they are likely to be severely affected by greater costs of information asymmetry, agency conflicts, costs associated with added layers of corporate bureaucracy, inefficient local management and workers, and other costs that may arise from cultural, political and economic differences with the host country in comparison with domestic investments.
- Previous studies (e.g. <u>Doukas and Travlos, 1988</u>; <u>Morck and Yeung, 1991</u>, <u>1992</u>; <u>Doukas, 1995</u>), however, address corporate expansion through the purchase of external growth opportunities such as international acquisitions.

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- 9. The construction of Tobin's q is described in the <u>Appendix</u>
- 10. Foreign sales/total sales ratios were collected from the Value Line Investment Survey.
- 11. Insider ownership information, included for reference, was also obtained from the Value Line Investment Survey.
- 12. To make sure that the four-digit SIC code of the company represents its core business, we also require that it accounts for 50% or more of its total sales. Other studies have routinely used only SIC codes to identify the core line of business of a corporation. Foreign investment activity based on the two-digit SIC code classification produces smaller diversification intensities.

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bidder returns at the announcement of conglomerate acquisitions in the late 1960s and early 1970s. <u>Ravenscraft and Scherer (1987)</u> record that conglomerate acquisitions during the 1960s were unsuccessful because of the post-acquisition poor performance and subsequent bust-up. Similar evidence is also reported by <u>Kaplan and Weisbach (1992)</u>. <u>Servaes (1996)</u> finds a negative relation between diversification and firm performance in the 1960s, and a negative but weaker relation in the 1970s.

17. Lummer and McConnell (1990) show that the formation of international joint ventures increases a firm's value by 0.4%. Finnerty et al. (1986), however, find insignificant announcement effects for 110 international joint ventures. Similarly, Lee and Wyatt (1990) report significantly negative stock price reactions for US-foreign joint venture announcements. The overall evidence is different from that reported in McConnell and Muscarella (1985), which

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- 21. An alternative approach to control for industry effects is to use an industry indicator variable (<u>Bradley et al., 1984</u>).
- 22. <u>Barber and Lyon (1997)</u> note that several biases (that is, new listing rebalancing, skewness, and a negative bias in continuously compounded returns) are introduced in long-term performance studies when the BHAR method is not used.
- 23. The BHARs are measured from day +3 to +500 (year 2) and +750 (year 3), respectively.

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- 27. The estimation of the Herfindahl index for the announcement year is estimated as the sum of the squared assets per business segment prior to the investment announcement plus the square of the new investment relative to the squared assets of all segments prior to the announcement plus the new investment. This method is different from the conventional Herfindahl index, which can be constructed only for year −1 using the Compustat segment data tape.
- 28. This theory was introduced by <u>Coase (1937)</u> and subsequently developed by <u>Caves (1971)</u>, <u>Dunning (1973)</u>, <u>Williamson (1975)</u>, and <u>Buckley and Casson (1976)</u>.

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significant bearing on the initial regressions, the reported results are based on regressions excluding these control variables. Only a very small fraction of firms is involved in subsequent foreign investments. Five per cent of the firms engage in foreign investment activity over the 2-year period and 7% in the 3-year period after the initial announcement. These results are available upon request.

- 34. Similar results, available upon request, are found when we use the other two measures of performance (EBITD/total assets and EBITD/market value of equity).
- 35. The consistency between the market's negative (positive) reaction to foreign investment announcements and the long-term operating decrease (increase)

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Appendix

Tobin's *q* estimates are obtained using the <u>Perfect and Wiles (1994)</u> method, which is a modified version of the <u>Lindenberg and Ross (1981)</u> algorithm for generating firm *q* values. The source of the data used to estimate *q* values is contained in the Compustat data tape, the Business Conditions Digest, and the Moody's Industrial Manuals. <u>Perfect and Wiles (1994)</u> determine the market value of the firm by the sum of:

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where *RC* is the replacement cost of the firm's assets; *Comval* is the year-end value of the firm's common stock; *Perfval* is the estimated year-end market value of the firm's preferred stock (that is, firm's total preferred dividends capitalized by the Standard and Poor's preferred stock yield index); *LTDebt* measures the value of the long-term debt with maturity greater than 1 year; and *STDebt* is the year-end book value of the firm's short-term debt with maturity less than 1 year.

The following formula is used to estimate the *LTDebt* variable:

$$\begin{split} & TDebt_t \\ &= SBond_t E_{j=0}^{n-2} f_{t,t-j} \\ & \times \Bigg\{ \left(\frac{R_{t-j}^A}{R_t^A} \right) [1 - (1 + R_t^A)^{-(n-j)}] + (1 + R_t^A)^{-(n-j)} \Bigg\}, \end{split}$$

where *SBond* is the year-end book value of the firm's long-term debt in year *t*: *f*

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