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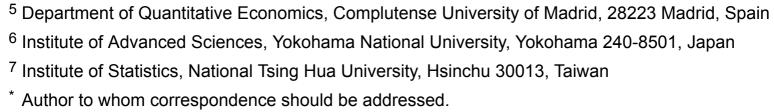
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Abstract

It is well known that there is an intrinsic link between the financial and energy sectors, which can be analysed through their spillover effects, which are measures of how the shocks to returns in different assets affect each other's subsequent volatility in both spot and futures (https://www.cookiebo ,Financial markeţs. derivatives, which are not only highly representative of the underlying on both the spot and futures markets, include Exchange Traded Funds (ETFs), a tradable spot index whose aim is to replicate the return of an underlying benchmark index. When ETF futures are not available to examine spillover effects, "generated regressors" are useful for constructing both financial ETF futures and energy ETF futures. The purpose of the paper is to investigate the co-volatility spillovers within and across the U.S. energy and financial sectors in both spot and futures markets, by using "generated regressors" and a multivariate conditional volatility model, namely diagonal BEKK. The daily data used are from 23 December 1998–22 April 2016. The dataset is analysed in its entirety and is also subdivided into three distinct subsets. The empirical results show there is a significant relationship between the financial ETF and energy ETF in the spot and futures markets. Therefore, financial and energy ETFs are suitable for constructing a financial portfolio from an optimal risk management perspective and also for dynamic hedging purposes.

Keywords: exchange traded funds (/search?q=exchange+traded+funds); financial and energy sectors (/search?q=financial+and+energy+sectors); co-volatility spillovers (/search?q=co-volatility spillovers); spot and futures prices (/search?q=spot+and+futures+prices); g erated regressors (/search?q=diagonal+BEKK)

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1. IMerokaténigan

The Global Financial Crisis (GFC) was not only unexpected and unpredicted, but also had a marked and sustained impact on the world economy, in general, and also on international financial markets. After the GFC had subsided, oil prices recovered and stabilized at a price between US\$90 and US\$110 per barrel (see **Figure 1**). This period of relative stability lasted from January 2011–June 2014. However, in mid-2014, oil prices nosedived from a high of US\$107.95 per barrel to a low of US\$26.19 per barrel on 11 February 2016.

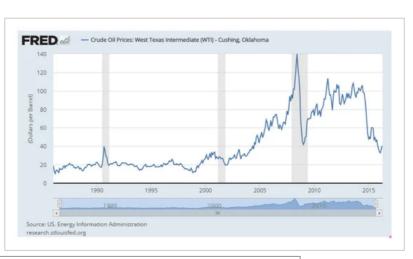


Figure 1. Crude oil prices: West Texas Intermediate (WTI) (1986–2016), Federal Reserve Economic Data.

According to a World Bank Report (**Baffes et al. 2015**), the plunge in oil prices was mainly driven by supply factors, namely the growth of unconventional oil production, such as Canadian oil sand and U.S. shale oil. In particular, spurred by the shale oil bo the production levels to over 11 million barrels in June 201 is Dehindupewatredebytheod (Sebat/) surpass

Saudi Arabia as the oil and natural gas liquids global production leader, as reported by the International Energy Agency (IEA) (Smith 2014, Bloomberg, 4 July).

Responding to the surge in unconventional oil production, at the 166th OPEC meeting held on 27 November 2014, OPEC decided not to curtail daily production, choosing instead to maintain a stable production of 30 million barrels per day, a policy enacted on 14 December 2011. This decision represented abandonment of OPEC's price targeting policy, with the trade-off of possibly maintaining their current market share. However, this course of action may well have led to persistently low oil prices.

increase reserves for losses in the oil and gas portfolio, banks have also tried to shrink the credit lines offered to energy companies, even as energy companies become more dependent on banking lans. This sentiment was echoed by Devi Aurora, a senior director at Standard & Poor's in New York, who was reported to have said (McLannahan and Gray 2016): "[Energy] Companies have a tendency to draw of the options dry up."

Such low oil prices have major ramifications on the banking sector. In addition to having to

Faced with the dual pressures of low oil prices and a compromised ability to generate cash flows, oil segmentes are increasingly in danger of defaulting on loans. As reported in the Wall Stree 'cournal, "Coming to the Oil Patch: Bad Loans to Outnumber the Good", 24 March 2016 (Olson et al. 2016):

Marketing
"Fifty-one North American oil-and-gas producers have already filed for bankruptcy since the start
of 2015, cases totaling \$17.4 billion in cumulative debt, according to law firm Haynes and Boone
LLP. That trails the number from September 2008 to December 2009 during the global financial
Show details >
crisis, when there were 62 filings, but is expected to grow: About 175 companies are at high risk
of not being able to meet loan covenants, according to Deloitte LLP."

From recent data, it is clear that oil price collapses of greater than 50% are not unprecedented events. For example, in 1986, there was a similar supply glut, which also led to a plunge in oil prices. In particular, that year marked OPEC's decision to revert its production target back to 30 million

barrels per day, ending a significant decline in oil production since the Iran-Iraq war in 1979. This reversion, combined with an influx of oil supply from Mexico and the North Sea, caused the price of oil to collapse from US\$26.53 per barrel on 6 January 1985 to US\$10.25 per barrel at its low point on 31

March 1986.

Around this time, the U.S. government attempted to stimulate the sluggish economy and guard

against deflation through several monetary and fiscal policies such as interest rate cuts. In spite of these measures, low oil prices persisted, thereby contributing to a global economic slowdown and a major downward correction in global financial markets on 19 October 1987. This day, known as Black Monday, saw the S&P 500 drop 20.4%, falling from 282.7 down to 225.06.

Another significant plunge in oil prices, this time of the order of 40%, occurred between October 1997 and March 1998 amid the Asian Financial Crisis. This crisis was propelled primarily by an

unexpected speculative attack on the Thai baht. The resulting drastic devaluation of the Thai currency not only wrought considerable damage to the East Asian economy, but also had a significant impact on global financial markets. The US Federal Reserve decided to bail out a well-known hedge fund,

Long Term Capital Management (LTCM), on 23 Septemb**ள்பி998/லெஸ்.ருdok iebotocoim ச்பா/முந்வ**hich lasted from 1997–1998, the global oil demand receded s**it stahindy powerdchoves oekishot/à** low of US\$10.82 per barrel on 10 December 1998.

The most dramatic example of a sudden oil price collapse occurred a decade later in the wake of the Global Financial Crisis (GFC). While there is no consensus on the exact starting and ending dates of the GFC, for the purposes of this paper, we consider the GFC to span the time period from 9

of the GFC, for the purposes of this paper, we consider the GFC to span the time period from 9 October 2007–9 March 2009, which corresponds to the S&P 500 dropping from a high of 1565.26 to a low of 672.88. Oil prices reached a historical high of US\$145.31 per barrel on 3 July 2008, but

tumbled to US\$30.28 per barrel just six months later, on 23 December 2009.

The GFC was spurred by a tsunami of financial chaos, including the housing bubble, which, in turn, led to an epidemic of defaults in subprime mortgages. Subsequently, banks and insurance companies sold trillions of dollars of Credit Default Swaps (CDSs), which not only involved subprime mortgage loans, but also many other financial instruments and institutions. This resulted in Lerman Brothers going bankrupt on 15 September 2008, and the U.S. Treasury deciding to bail out AIG in the

energy prices (Van Vactor 2009).

In light of the preceding discussion, it is clear that there is an intrinsic link between the financial and the sectors. One way to analyse the link between two or more sectors is by analy gotheir spillover effects, which are measures of how the shocks to returns in different assets affect each

same month. The GFC led to a dramatic diminution in the global oil demand and, in turn, tumbling

Marketing In conducting spillover effects analysis, an important consideration is the choice of indices used to represent the assets or sectors under comparison. One reasonable selection of measures to examine volatility spillovers between the energy and financial sectors is the Energy Select Sector index (Ticker: IXE) and the Financial Select Sector index (Ticker: IXM). Both of these are sub-indices

of the S&P500, reflecting the overall economic condition of their respective sectors. One shortcoming of using these indices, however, is the fact that they are not tradable and hence may be of little practical use to investors.

One way to overcome this drawback is by employing derivatives of the IXE and IXM indices, as

opposed to the indices themselves. Financial derivatives, which are not only highly representative of the underlying indices, but can also be traded on both the spot and futures markets, include Exchange Traded Funds (ETFs), otherwise known as implied tradable spot prices. Another financial derivative that has not yet been considered in practice, primarily as it typically does not exist in many

financial markets, but may well have practical importance, is ETF futures.

Krause and Tse (2013) argue that ETFs are now an important source of information dissemination in equity markets and provide evidence regarding uni-directional price discovery and bi-

dissemination in equity markets and provide evidence regarding uni-directional price discovery and bidirectional volatility spillovers in ETFs. The empirical findings are relevant to market participants and Canadian market regulators as circuit-breakers in Canada are tied to U.S. market conditions.

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other's subsequent volatility in both spot and futures markets.

Lau et al. (2017) investigate the spillovers and volatility transmission among white precious metals and gold, oil and global equity. Relying on ETFs, they analyse return spillovers using an

EGARCH model and frequency dynamics (see (McAleer and Hafner 2014a, 2014b; Chang and McAleer 2017a) for some serious issues associated with EGARCH). The empirical results highlight

the role of gold ETFs as the most influential market.

With respect to multi-objective portfolio optimization (https://wwwiwacoookietaoketomSen/ikh(2008) considers a three-stage lexicographic approach for m**istibehind-powerted-bytiፍααklebot**ሩ). mixed integer programming, while **Sawik (2012a**) analyses bi-criteria portfolio optimization models with

percentile and symmetric risk measures by mathematical programming, and Sawik (2012b) examines

a downside risk approach for multi-objective portfolio optimization. For the reasons specified above, in order to examine the relationship between the energy and

financial sectors, we will examine not only IXE and IXM, but also ETFs and ETF futures in conducting spillover effects analysis within and across these two sectors. In particular, for both the energy and financial sectors, we will select one index (namely, IXE or IXM), one ETF and construct one ETF

futures from which to analyse all 15 possible pairwise combinations of spillover effects. The list of indices, ETFs and ETF futures that we will use in the empirical analysis is as follows: Financial Select Sector Index (IXM), Energy Select Sector Index (IXE), Financial Select Sector SPDR Fund (XLF),

Energy Select Sector SPDR Fund (XLE), Financial ETF futures (XLFf) and Energy ETF for the sector SPDR Fund (XLE), Financial ETF futures (XLEf). An important point to clarify is that, despite the delisting of ETF futures on 1 March 2011, due to low**Ptratingnyon**tme, our analysis will include up-to-date ETF futures data from each sectc. This is

made possible by the use of "generated regressors" to construct both financial ETF futures and energy ETE futures. Further details on this methodological approach are discussed in Section An Exchange Traded Fund (ETF) is a tradable spot index whose aim is to replicate the return of an underlying benchmark index. For instance, SPDR $^{
m @}$ S&P 500 $^{
m @}$ ETF, issued by State Stree $^{
m t}$ Bank &

and decreasing the levels of risk and volatility. Furthermore, unlike actively-managed mutual funds, most ETF managers take a passive management style and collect lower managing fees. Whereas mutual funds are limited to trades based on end-of-day prices, ETFs are traded like stocks.

Marketing
Trust Company, tracks the performance of the S&P 500 Index. In contrast to investing in a single stock, ETFs invest in a basket of stocks or commodities, thereby diversifying the non-systematic risk

In addition to the points listed above, ETFs have the following additional advantages over traditional mutual funds: (i) ETFs offer greater transparency compared with mutual funds in the sense that ETFs are required

to reveal their holdings data on a daily basis, whereas mutual funds are mandated only to disclose holdings data on a quarterly basis. (ii) ETFs are more flexible than mutual funds because investors can short sell them when they are bearish on the market. Although short selling may be considered risky compared with conventional

investing, it can be a useful strategy if executed by savvy investors when the market is overvalued. As discussed above, the purpose of this paper is to investigate spillover effects within and across the energy and financial sectors in terms of both the U.S. spot and futures markets by applying

indices, ETF and ETF futures. For the empirical analysis, we select two indices and two ETFs and Loading [MathJax]/jax/output/HTML-CSS/fonts/Gyre-Pagella/Size2/Regular/Main.js

2. Brief Literature Review

The literature on the use of ETFs and testing for co-volatility spillovers is rather sparse. Chang et ai. (2015) conducted a comprehensive review of the literature related to co-volatility spillovers

generate two ETF futures from which to analyse all 15 possible pairwise combinations of spillover effects. Specifically, the list of variables we use is as follows: Financial Select Sector Index (IXM), Energy Select Sector Index (IXE), Financial Select Sector SPDR Fund (XLF), Energy Select Sector SPDR Fund (XLE), Financial ETF futures (XLFf) and Energy ETF futures (XLEf). In order to carry out this analysis, the techniques to be used are generated regressors and the multivariate conditional

volatility diagonal BEKK model. The empirical result will b**é<u>htitps://swdviw.greakéebletadamSer/tidna 5</u>.**

is reviewed. In **Section 3**, the empirical models are presented, and the data are discussed in **Section 4**. In **Section 5**, the empirical results are analysed, and some concluding comments are given in

The remainder of the paper is organized as follows. Is section 20 the paper is organized as follows.

between energy markets and agricultural commodities. One of the major findings of their review paper was that most researchers fail to employ valid statistical techniques in testing for spillover effects. Multivariate conditional volatility models, namely BEKK and DCC, have typically been used to test for spillovers between energy and agricultural markets. However, these models are the

problematic in and of themselves (in the case of DCC), or have been used incorrectly and mistaglically dies the case of BEKK).

Specifically, the DCC model lacks regularity conditions, while a serious technical deficiency related to estimating the full BEKK and DCC models through Quasi-Maximum Likelihood Estimates (QMLE) is the absence of any asymptotic properties. In contrast, the diagonal BEKK conditional

volatility model possesses both regularity conditions and associated asymptotic properties. For these

volatility spillovers for bio-ethanol, sugarcane and corn spot and futures prices, while this paper also

realtankethang et al. (2018) applied the diagonal BEKK conditional volatility model

applies the diagonal BEKK model in testing the volatility spillover effects within and across the U.S. financial and energy markets. In this context, **Chang and McAleer** (**2017b**) degelocating ply test for causality in volatility, which can be used for testing volatility spillovers.

As described above, an Exchange Traded Fund (ETF) is a tradable asset whose aim is to track

an underlying index representing the economic condition of an entire sector. Thus, ETFs have great value to investors as they facilitate a systematic reduction in risk within a trading portfolio. **Chang and Ke** (2014) applied ETFs in the U.S. energy sector to investigate the causality between flows and

returns through the Vector-Autoregressive (VAR) model to test four hypotheses, namely the price pressure, information, feedback trading and smoothing hypotheses. One noteworthy aspect of their methodology was the fact that they analysed not just the entire sample period, but also divided the data into three sub-periods, namely before, during and after the Global Financial Crisis (GFC), a

methodology also used by **McAleer et al.** (2013). The use of the three sub-periods will also be considered in the paper.

Chen and Huang (2010) used ETFs to examine volatility spillovers, albeit in a rudimentary

Chen and Huang (2010) used ETFs to examine volatility spillovers, albeit in a rudimentary manner, between an ETF and its underlying stock index in nine different countries. They used the GARCH-ARMA and EGARCH-ARMA models and found that there were volatility spillover effects for Loading [MathJax]/jax/output/HTML-CSS/fonts/Gyre-Pagella/Size2/Regular/Main.js

One paper that used the diagonal BEKK model to examine ETFs was by **Chang et al.** (2016). The authors investigated the causality and spillover effettpsb#tweencolookiebotsistimgenfwhitetrent moving average processes, and ETF returns by using sethind progressite. (4) Els and diagonal BEKK models. The empirical results show that daily VIX returns have: (1) significant negative effects on European ETF returns in the short run; (2) stronger significant effects on single market ETF returns than on European ETF returns; and (3) lower impacts on the European ETF

In some financial research contexts, it may be necessary or advantageous to generate a new

the stock index and ETF. Unfortunately, as in the case of estimating the full BEKK and DCC models through Quasi-Maximum Likelihood Estimation (QMLE) methods, EGARCH has no known regularity conditions (see **McAleer and Hafner 2014a**), and the statistical properties of the estimated

parameters are not available except by assumption (see Martinet and McAleer 2018).

index representing a certain sector that may be of interest. One way in which this may be performed is through the use of generated variables. **Chang** (2015) applied generated variables to develop a daily Tourism Financial Conditions Index (TFCI), based on nominal exchange rates, interest rates and a tourism industry stock index that is listed on the Taiwan Stock Exchange. The empirical results indicated that the generated TFCI was accurately estimated through the estimated conditional means of the tourism stock index returns. More recently, **Chang et al.** (2017) used a similar approach to obtain a monthly Tourism Financial Conditions Index (TFCI) based on factor analysis and macroeconomic variables to capture general economic activity. The use of market returns on the tourism sector, as compared with the general activity

As described above, the paper is interested in the co-volatility spillover effects across and within the financial and energy sectors in both the spot and futures markets. While energy and financial Marketing indices and ETFs are already available to analyse spot markets, it is necessary to use generated variables to construct ETF futures to analyse futures markets.

The paper combines several of the elements reviewed above to create a novel methodology to

of economic variables on tourism stocks using TFCI, is shown to provide an exaggerated and

practical use to investors. In particular, we use the diagonal BEKK model, which, as mentioned above, has valid asymptotic and regularity properties as compared with the full BEKK and DCC models, in order to test for spillovers within and across the financial spot (indices and ETFs) and futures (ETF futures via generated regressors) markets. This analysis is conducted for four time periods, namely before-GFC, during-GFC, after-GFC and the entire sample period.

test for spillover effects in a statistically valid and comprehensive way that can be of immense

3. Methodology

returns than on S&P500 returns.

The primary purpose of this paper is to test volatility spillover effects among ETF and ETF futures in the financial and energy sectors. In the previous literature, a great deal of confusion has arisen about how spillover effects should be tested, with published academic papers often using

questionable methodologies. Indeed, many so-called tests of spillovers are not, in fact, tests of spillovers at all. The following section presents three novel tests of spillovers, namely full volatility

excessively volatile explanation of tourism financial conditions.

et al. (2015). Tests of spillovers require the estimation of a multivariate volatility model, with appropriate regularity conditions and asymptotic properties of the Quasi Maximum Likelihood Estimation (QMLE) of the associated parameters underlying the conditional mean and conditional variance. As the first

spillovers, full co-volatility spillovers and partial co-volatility spillovers. For further details, see Chang

step of the estimation of multivariate conditional volatility relations in the estimation of multivariate conditions in the estimation of the estim conditional volatility models, an appropriate and widely-us**isdbethiadate wered்dbed codsittly ot**/ødel will be discussed below. This section is organized as follows:

(1) a brief discussion of the most widely-used univariate conditional volatility model;

- (2) the definition of three novel spillover effects; (3) a discussion of the most widely-used multivariate model of conditional volatility.

In order to accommodate volatility spillover effects, alternative multivariate volatility models of the

conditional covariances are available. Examples of such multivariate models include: (1) the diagonal

model of Bollerslev et al. (1988); (2) the vech and diagonal vech models of Engle and Kroner (1995); (3) the Baba et al. (1985) (BEKK) multivariate GARCH model (see also (Engle and Kroner

1995) gestative Constant Conditional Correlation (CCC) (specifically, multiple univariate rathe multivariate) GARCH model of Bollerslev (1990) (5) Ling and McAleer's (2003) vector ARMA-

GARCH (VARMA-GARCH) model; (6) the VARMA-asymmetric GARCH (VARMA- AGARCH) model of McAleer et al. (2009); (7) Engle's (2002) Dynamic Conditional Correlation (technically, dynamic

conditional covariance rather than correlation model) (DCC) model; and (8) Tse and Tsui's (2002) Var**ัรโลรู้ เอีย**์คื**อ**itional Correlation (VCC) model. For further details on most of these multivariat⊾ nodels, see, for example, McAleer (2005).

conditional mean returns shocks. For this reason, the most widely-used univariate conditional volatility model, namely GARCH, will be presented briefly, followed by the most widely-estimated multivariate conditional covariance model, namely diagonal BEKK. Show details > Consider the conditional mean of financial returns as follows:

where the returns, , represent the log-difference in financial commodity or agricultural prices, , is the information set at time t – 1 and is a conditionally heteroskedastic returns shock. In order to derive

(1)

Marketistastep in estimating multivariate models is to obtain the standardized shocks

conditional volatility specifications, it is necessary to specify the stochastic processes underlying the returns shocks, . 3.1. Univariate Conditional Volatility Models

Alternative univariate conditional volatility models are of interest in single index models to describe individual financial assets and markets. Univariate conditional volatilities can also be used to standardize the conditional covariances in alternative multivariate conditional volatility models to estimate conditional correlations, which are particularly useful in developing dynamic hedging

strategies. The most popular univariate conditional volatility model is discussed below, together with the associated regularity conditions, as well as the conditions underlying the asymptotic properties of Loading [MathJax]/jax/output/HTML-CSS/fonts/Gyre-Pagella/Size2/Regular/Main.js

consistency and asymptotic normality. A deeper discussion of the material presented in this section is available in, for example, **Chang et al.** (2018).

Random coefficient autoregressive process and GARCH

Consider the random coefficient autoregressive process of order one:

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where and is the standardized residual.

Tsay (1987) derived the ARCH(1) model of Engle (1982) from Equation (2) as:

positive, as they are the unconditional variances of two separate stochastic processes.

where is conditional volatility and is the information set available at time t-1. The use of an infinite lag length for the random coefficient autoregressive process in Equation (2), with appropriate geometric restrictions (or stability conditions) on the random coefficients, leads to the GARCH model

The QMLE of the parameters of ARCH and GARCH have been shown to be consistent and asymptotically normal in several papers. For example, **Ling and McAleer** (2003) showed that the QMLE for GARCH(p,q) is consistent if the second moment is finite. Moreover, a weak sufficient log-molecular molecular for the QMLE of GARCH(1,1) to be consistent and asymptotically normal is given

of Bollerslev (1986). From the specification of Equation (2), it is clear that both and should be

The more restrictive second moment condition, namely, is much easier to check in practice.

Statistics al, the proofs of the asymptotic properties follow from the fact that ARCH and ARCH can be derived from a random coefficient autoregressive process (see McAleer et al. (2008) for a

which is not easy to check in practice as it involves two unknown parameters and a random variable.

general proof of multivariate models that are based on proving that they satisfy the gularity conditions given in **Jeantheau** (1998) for consistency).

3.2. Multivariate Conditional Volatility Models

vectors, where is the number of financial assets.

The multivariate extension of univariate GARCH is given as variations of the Betail model in Baba et al. (1985) and Engle and Kroner (1995). In order to establish volatility spillovers in a multivariate framework, it is useful to define the multivariate extension of the relationship between the returns shocks and the standardized residuals, that is . The multivariate extension of Equation (1), namely), can remain unchanged by assuming that the three components are now, respectively,

The multivariate definition of the relationship between and is:

where is a diagonal matrix comprising the univariate conditional volatilities. Define the conditional covariance matrix of as. As the vector, , is assumed to be independently and identically distributed (iid) for all elements, the conditional correlation matrix of , which is equivalent to the conditional correlation matrix of , is given by . Therefore, the conditional expectation of (4) is defined as:

(5)

(3)

(7)

(8)

(9)

(10)

Equation (5) is useful if a model of is available for purposes of estimating, whereas Equation (6) is useful if a model of is available for purposes of estimating.

Equation (5) is convenient for a discussion of volatility spillover effects, while both Equations (5) and (6) are instructive for a discussion of asymptotic properties. As the elements of are consistent and asymptotically normal, the consistency of in (5) dependence of in (5) dependence of in (6) depends on consistent estimations of the definitions given in Equations (5) and (6).

3.3. Full and Partial Volatility and Co-Volatility Spillovers

Volatility spillovers are defined in **Chang et al.** (2015) as the delayed effect of a returns shock in one asset on the subsequent volatility or co-volatility in another asset. Therefore, a model relating to returns shocks is essential, and this will be addressed in the following sub-section. Spillovers can be defined in terms of full volatility spillovers and full co-volatility spillovers, as well as partial co-volatility spillovers, as follows:

(1)Full volatility spillovers:

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(2)Full co-volatility spillovers:

Preferences

(3) Partial co-volatility spillovers:

Statistics

where is returns shocks and is the conditional covariance matrix of .

hedging. Full volatility spillovers occur when the returns shock from financial asset k affects the volatility of a different financial asset i. Full co-volatility spillovers occur when the returns shock from **Show details** financial asset k affects the co-volatility between two different financial assets, i and j. Partial co-volatility spillovers occur when the returns shock from financial asset k affects the co-volatility

Volatility spillovers in the spot and derivatives markets are crucial for purposes of dynamic

between two financial assets, i and j, one of which can be asset k. When m = 2, only Options (1) and (3) are possible as full co-volatility spillovers depend on the existence of a third financial asset.

As mentioned above, spillovers require a model that relates the conditional volatility matrix, , to a matrix of delayed returns shocks. The most frequently-used models of multivariate conditional covariance are alternative specifications of the BEKK model, with appropriate parametric restrictions, which will be considered below.

3.4. Diagonal and Scalar BEKK

The vector random coefficient autoregressive process of order one is the multivariate extension of Equation (2) and is given as:

where and are vectors, is an matrix of random coefficients and:, .

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Technically, a vectorization of a full (that is, non-diagonal or non-scalar) matrix A to vec A can have a dimension as high as , whereas vectorization of a symmetric matrix A to vec A, starting with

the diagonal elements for stacking the matrix, can have a dimension as low as .

In the case where A is either a diagonal matrix or the special case of a scalar matrix, , **McAleer**

et al. (2008) showed that the multivariate extension of GARCH(1,1) from Equation (10), incorporating an infinite geometric lag in terms of the returns shocks (histogical was with each interportation).

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where *A* and *B* are both either diagonal or scalar matrices. The matrix *A* is crucial in the interpretation of symmetric and asymmetric weights attached to the returns shocks, as well as the subsequent

analysis of spillover effects.

McAleer et al. (2008) showed that the QMLE of the parameters of the diagonal or scalar BEKK models were consistent and asymptotically normal, so that standard statistical inference on testing

estimated consistently.

In terms of volatility spillovers, as the off-diagonal terms in the second term on the right-hand side of Equation (11), have typical (*i,j*) elements, there are no full volatility or full co-volatility spillovers.

hypotheses is valid. Moreover, as in (11) can be estimated consistently, in Equation (6) can also be

Ho**₩ever, Spati**tial co-volatility spillovers are not only possible, but they can also be tested using alic statistical procedures. 3.5Pi**retermentary**

Without actually deriving the model from an appropriate stochastic process, Baba et al. (1985)

3.3F Indexiguilar, 3 lavallaru allu Full BEKK

multivariate model is the same as the specification in Equation (11), namely:

Marketing

(12)

except that A and B are full, Hadamard or triangular matrices, rather than diagonal or scalar matrices,

triangular and Hadamard (element-by-element multiplication) BEKK models. The specification of the

andstandstand Kroner (1995) considered the full BEKK model, as well as the special

as in (11).

Show details >

Although it is possible to examine spillover effects using each of these models, it is not possible

to test or analyse spillover effects as the QMLE of the parameters in Equation (12) have no known asymptotic properties.

Although estimation of the full, Hadamard and triangular BEKK models is available in some standard econometric and statistical software packages, it is not clear how the likelihood functions might be determined. Moreover, the so-called "curse of dimensionality", whereby the number of parameters to be estimated is excessively large, makes convergence of any estimation algorithm

somewhat problematic. This is in sharp contrast to a number of published papers in the literature, whereby volatility spillovers have been tested incorrectly based on the off-diagonal terms in the matrix

3.6. Generated Regressors

A in Equation (12).

One of the primary purposes of the paper is to investigate the spillover effects within and across the energy and financial sectors for both U.S. spot and futures market by applying indices, ETF and

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necessary to use generated variables to construct ETF futures for futures markets. The generated ETF futures proposed in the paper focus on economic activities related to the financial and energy industries, respectively. The three components of the Financial ETF futures (XLFf), each of which can

ETF futures. While energy and financial indices and ETFs are already available for spot markets, it is

be constructed from data downloaded from Bloomberg or Yahoo Finance, are as follows:

(https://www.cookiebot.com/en/whatis-behind-powered-by-cookiebot/)

(2) Generic 1st S&P 500 index futures (SP1); and

(3)Generic 1st FTSE 100 index futures (Z1).
The other three components of the energy ETF futures (XLEf), each of which can be constructed

(3)Generic 1st Natural Gas futures (NG1).

from data downloaded from Bloomberg or Yahoo Finance, are as follows:

(1) Energy Select Sector SPDR Fund (XLE);
(2) Generic 1st Crude Oil WTI futures (CL1); and

The ETF futures discussed above are based on estimation of a regression model, which may be referred to as the generating model. The model-based weights for the components of financial ETF futures are estimated by OLS. The traditional method of examining the statistical properties of generated variables, and more specifically generated regressors, uses variables typically stationary. In empirical finance, the variables considered can be nancial

returns, in which the variables are typically stationary, or financial stock prices, where the variables

are typically non-stationary.

Statistics

The specific model that is used to generate ETF futures is based on financial price variables, all of which are non-stationary. Consequently, there would seem to be no known optimality properties for the MOLSE interprets of ETF futures. For this reason, the generated variable is interprets as an estimate of ETF, with no optimal statistical properties claimed for the estimated parameters in the

generating model. In comparison, where the variables are stationary, Ordinary Least Squares (OLS)

can be shown to be efficient (see, for example, (**McAleer and McKenzie 1991**; **ᢂbAtestet1993**; **Fiebig** et al. 1992)). For the purposes of determining whether the generated ETF futures are a reasonable

construction of the latent variable, will be used as a statistical indicator.

The models to be estimated below are linear in the variables, with the appropriate weights to be

estimated empirically. Accordingly, is defined as:

(13)

where c denotes the constant term and denotes the shocks to XLFf, which need not be independently or identically distributed, especially for daily data. The parameters θ_1 , θ_2 and θ_3 are the weights

attached to one-period lagged financial ETF, Generic 1st S&P 500 index futures and Generic FTSE

100 index futures, respectively.
As XLFf is a latent variable, it is necessary to link XLFf to observable data. The latent variable is defined as being the conditional mean of an observable variable, namely the Financial Select Sector

SPDR Fund (XLF), which is a tradable spot index, reflecting the financial select index that is listed on the NYSE, as follows:

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(15)

(16)

(17)

where XLF is observed, XLFf is latent and the measurement error in XLF is denoted by , which need not be independently or identically distributed, especially for daily data.

Given the zero mean assumption for , the means of XLF and XLFf will be identical, as will their estimates. Using Equations (13) and (14), the empirical model for estimating the weights for XLF is (https://www.cookiebot.com/en/whatgiven as:

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where , which should be distinguished from the return shocks, , in Equations (1) and (4) above, need not be independently or identically distributed, especially for daily data.

The parameters in Equation (15) can be estimated by OLS or QMLE, depending on the

specification of the conditional volatility of , to yield estimates of XLF, if SP1 and Z1 are stationary. As XLF is a non-stationary price, there is no reason to expect the combined error, , to be conditionally heteroskedastic. Alternatively, Instrumental Variables (IV) or the Generalized Method of Moments (GMM) can be used to estimate the parameters in Equation (15) to obtain an estimate of XLF and, hence, also an estimate of the latent variable, XLFf, although finding suitable instruments can be

est**inates of t**he parameters are desired, and if statistical inference is intended for the estimate.. As we are interested only in the fitted values of ETF to generate ETF futures, namely XLF to obtain XLFf, the peaternative methods are eschewed in favour of the Ordinary Least Squares (OLS) esti view of the definition in Equation (14), the estimates of XLF will also provide estimates of the latent

Cointegration can also be used to estimate the parameters in Equation (15), but only if consistent

XLFf. tatistics Similar logic to the above applies to the energy case. XLEf is defined as follows:

Marketing where XLE is observed, XLEf is latent and the measurement error in XLE is denoted by , which need not be independently or identically distributed, especially for daily data

Given the zero mean assumption for , the means of XLE and XLEf will be identical, as will their estimates. Using Equations (13) and (14), the empirical model for estimating the weights for XLE is given as:

where need not be independently or identically distributed, especially for daily data.

As there would seem to be no known optimality properties for the OLS estimates of ETF futures, the OLS estimates of XLE will be used to estimate XLEf, though no optimality properties are claimed for the generated XLE futures.

4. Data and Variables

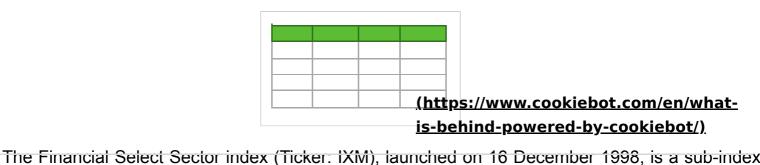
problematic when daily data are used.

As shown in **Table 1**, we choose the following indices, ETFs and ETF futures for the empirical analysis: Financial Select Sector Index (IXM), Energy Select Sector Index (IXE), Financial Select Sector SPDR Fund (XLF), Energy Select Sector SPDR Fund (XLE), Financial ETF futures (XLFf) and

Energy ETF futures (XLEf).

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 Table 1. Data description.



Industry Classification Standard (GICS®). The index represents the performance of the U.S. financial industry. Components of the Financial Select Sector are weighted by their float-adjusted market capitalization, and the Select Sector Indices are rebalanced quarterly. The three largest constituents of the financial sector are Berkshire Hathaway B, Wells Fargo & Co and JP Morgan Chase & Co. The

related ETF tracking IXM is the Financial Select Sector SPDR Fund (Ticker: XLF), as listed on the

New York Stock Exchange.

of S&P500 comprising 92 financial-related S&P 500 stocks. The classification is based on the Global

Correspondingly, the Energy Select Sector index (Ticker: IXE), launched on 16 December 1998, is a sub-index of S&P500 comprised of 38 energy-related stocks of the S&P 500. The classification is based on the S&P 500 industry Classification Standard (GICS®). This index represents the performance of the U.S. energy industry. Components of the Energy Select Sector are weighted by their float-

adjusted market capitalization, and the Select Sector Indices are rebalanced quarterly. The related ETFs tracking IXE is the Energy Select Sector SPDR Fund (Ticker: XLE), as listed on the New York Stock Exchange.

States Fice notical Select Sector SPDR® Fund (Ticker: XLF), issued by SSGA Funds Man ement,

Inc. and listed on the New York Stock Exchange since 16 December 1998, is the most representative financial ETF, with the largest total assets and average trading volume in the financial sector. This

ETF seeks to replicate the performance of the Financial Select Sector Index. As of 31 May 2016, the industry allocation of XLF consisted of banks (34.47%), Real Estate Investment Trusts (REITs) (18.30%), insurance (16.83%), diversified financial services (13.04%), capital markets (12.01%), consumer finance (4.92%), real estate management and development (0.29%) and unassigned (0.10%). The top three holdings of XLF are Berkshire Hathaway Inc. Class B (8.84%), JPMorgan Chase & Co. (8.04%) and Wells Fargo & Company (7.87%).

Correspondingly, the Energy Select Sector SPDR® Fund (Ticker: XLE), issued by SSGA Funds Management, Inc. and listed on the New York Stock Exchange since 16 December 1998, is the most representative energy ETF, with the largest total assets and average trading volume in the energy sector. This ETF seeks to replicate the performance of the Energy Select Sector Index. As of 31 May 2016, the industry allocation of XLE consisted of oil gas and consumable fuels (83.19%), energy

equipment and services (16.66%) and unassigned (0.15%). The top three holdings of XLE are Exxon

Mobil Corporation (18.85%), Chevron Corporation (14.68%) and Schlumberger NV (8.37%).

The financial ETF futures (XLFf) was generated from the Financial Select Sector SPDR® Fund (XLF), Generic 1st S&P 500 index futures (Bloomberg ticker: SP1) and Generic 1st FTSE 100 index

futures (Bloomberg ticker: Z1). The Generic 1st S&P 500 index futures is the continuous contract constructed by the front-month futures contract of S&P 500 index futures (Ticker: SPX), the latter having been an tendentially by the Generic

1st FTSE 100 index futures is the continuous contract constructed by front-month futures contract of FTSE 100 index futures, the latter having been launched by the London International Financial Futures and Options Exchange (LIFFE) in 1984.

Estimation of XLFf using generated regressors via the software R is shown in Equation (18):

1st S&P 500 index futures, Z1 is Generic 1st FTSE 100 index futures, and *t*-ratios are shown in parentheses. As stated previously, the *t*-ratios do not have the standard asymptotic normal distribution as the variables are non-stationary, but the extremely high value of suggests that the generated variable is a useful construction of the latent variable.

variable is a useful construction of the latent variable.

The Energy ETF futures (XLEf) are generated from the Energy Select Sector SPDR[®] Fund (XLE), Crude Oil futures (CL1) and Natural Gas futures (NG1). The Generic 1st Crude Oil futures is the continuous contract constructed by the front-month futures contract of Crude Oil WTI futures (Ticker: CL), listed in the New York Mercantile Exchange (NYMEX). The Generic 1st Natural Gas futures is the continuous contract constructed by the front-month futures contract of Natural Gas futures (Ticker: NG) listed in the New York Mercantile Exchange (NYMEX).

Necessary Estimation of XLEf using generated regressors via the software R is given in Equation (19).

```
XLEf_t = 0.0632 + 0.9989XLE_{t-1} + 0.00024CL1_{t-1} - 0.00325NG1_{t-1} (19)

Preferendes (875.35) (0.271) (-0.543)

\bar{R}^2 = 0.9986

where XLEf is Energy ETF futures, XLE is Energy Select Sector SPDR® Fund, CL1 is Generic 1st
```

Crude to the State of the state of the state of the state of suggests that the standard asymptotic normal distribution as the realist as the standard asymptotic normal distribution as the realist are non-stationary, but the extremely high value of suggests that the standard asymptotic normal distribution as the realist as a useful construction of the latent variable.

Daily data for the financial select sector index, energy select sector index, financial ETF, energy ETF and the constituents of the financial ETF futures and energy ETF futures (pamely, Generic 1st

S&P 500 index futures, Generic 1st FTSE 500 index futures, Generic 1st Crude Oil futures and Generic 1st Natural Gas futures) were downloaded from Bloomberg or Yahoo Finance. In the case of a national holiday, the missing value is replaced by the value of the previous day. ETF fund returns

are calculated by taking the log difference of adjusted prices and multiplying by 100, that is. The relevant descriptive statistics are shown in **Table 2**, implying that the returns of all variables are not normal. The Augmented Dickey–Fuller (ADF) and PP (Phillips–Perron) test for unit roots are shown in **Table 3**. The unit roots tests indicate that the returns of all variables are stationary.

Table 2. Descriptive statistics (22 December 1998–22 April 2016).

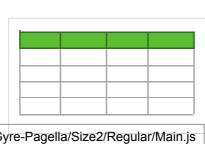
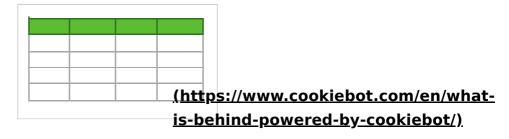


Table 3. Unit Root Tests. ADF, Augmented Dickey-Fuller; PP, Phillips-Perron.



The empirical analysis was conducted in its entirety and also subdivided into three sub-periods, namely (i) before-GFC, from 22 December 1998–8 October 2007; (ii) during-GFC, from 9 October 2007–9 March 2009; (iii) after-GFC, from 10 March 2009–22 April 2016; (iv) all (full sample), from 22 December 1998–22 April 2016. The numbers of observations for each period are 2292, 370, 1859 and 4521, respectively.

5. Empirical Results for Co-Volatility Spillovers

5.1. Hypothesis Testing of Co-Volatility Spillovers

This paper uses the diagonal BEKK model, in which the co-volatility spillover effects are a function of the diagonal elements of matrix A and the returns shocks of asset i at time t - 1. A **Necessary** rejection of the null hypothesis H₀, as shown in the definition of the test of co-volatility spillover enects in **Section 3**, indicates the significance of the co-volatility spillovers from the returns shocks of asset j at the terms of the co-volatility between assets i and j at time t.

In the empirical analysis, we selected two indices and two ETFs and generated two ETF futures, from the two transfers and two effects base on the multivariate diagonal BEKK model, specifically the co-volatility spillovers for all cases in which the estimates of A in the diagonal BEKK model are significant. The diagonal BEKK model shown in Equation (11) was estimated by QMLE using the econometric software package EViews 8.

The list of variables used is as follows: Financial Select Sector Index (IXM), Energy Select Sector Index (IXE), Financial Select Sector SPDR Fund (XLF), Energy Select Sector SPDR Fund (XLE), Financial ETF futures (XLFf) and Energy ETF futures (XLEf).

5.2. Calculating Average Co-Volatility Spillovers

Table 4 shows the estimates of the diagonal elements of *A* in the diagonal BEKK model for each pairwise comparison analysed (as described below), while **Table 5** shows the mean returns shocks for each asset, both for the entire time period and for each of the three sub-periods. **Table 6** shows the mean co-volatility spillovers, which are calculated by applying the definition of the co-volatility spillover effects discussed in **Section 3**.

Table 4. Estimation of diagonal elements of A in BEKK. GFC, Global Financial Crisis.

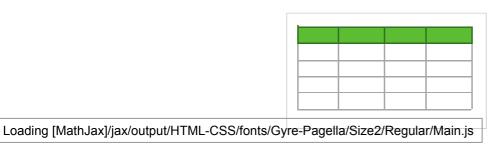
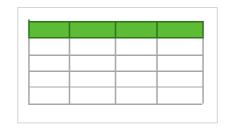


Table 5. Mean return shocks.



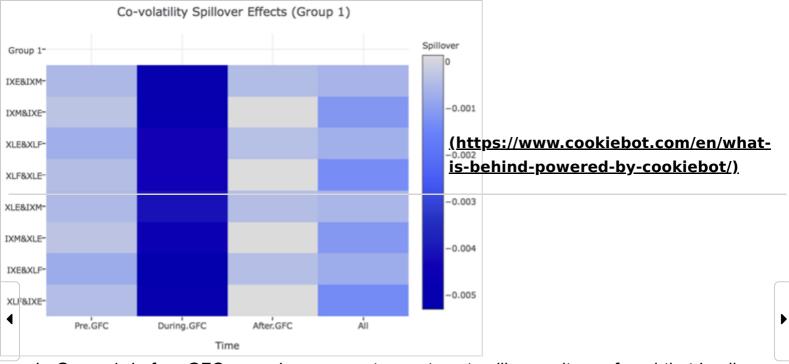
Table 6. Mean co-volatility spillovers.



As can be seen in **Table 6** and the explanation below, the data were separated into five groups, which will be described in detail below.

- Group 1: Cross-sector spot-spot spillover effects, specifically the spillover effects between each of the pairs; (a) financial index and energy index, (b) financial ETF and energy ETF, (c) financial index and energy energy ETF and (d) energy index and financial ETF.
- Group 2: Cross-sector futures-futures spillover effects, specifically the spillover effects between **Preferences**(a) financial ETF futures and energy ETF futures.
- Group 3: Cross-sector spot-futures spillover effects, specifically the spillover effects between each statisticspairs: (a) financial index and energy ETF futures, (b) financial ETF and energy ETF futures, (c) energy index and financial ETF futures and (d) energy ETF and financial ETF futures.
- Group 4: Within-sector spot-spot spillover effects, specifically the spillover effects between (a) **Marketing** financial index and financial ETF and (b) energy index and energy ETF.
- Group 5: Within-sector spot-futures spillover effects, specifically the spillover effects between each of the pairs: (a) financial index and financial ETF futures, (b) financial ETF and financial ETF **Show details** tutures, (c) energy index and energy ETF futures and (d) energy ETF and energy ETF futures.

The following paragraphs describe the average co-volatility spillover effects for each of the five groups mentioned above and also across each of the four time periods, namely "before- GFC", "during-GFC", "after-GFC", and "all".

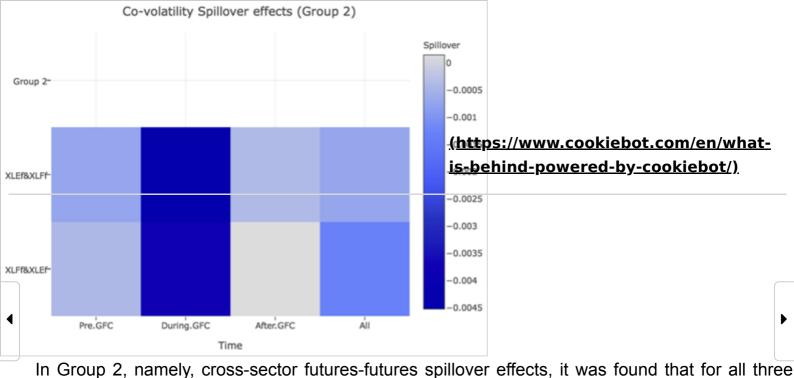


In Group 1, before-GFC, namely cross-sector spot-spot spillovers, it was found that in all cases, co-volatility spillovers were statistically significant and negative. For each of the four pairs, the magnitude of the spillovers of the financial spot asset, namely IXM (Financial Select Sector Index) or XLF (financial ETF), on subsequent co-volatility between itself and its corresponding energy spot asset, namely IXE (Energy Select Sector Index) or XLE (energy ETF), was numerically greate. Another spillovers of the energy spot asset on the same subsequent co-volatility pair.

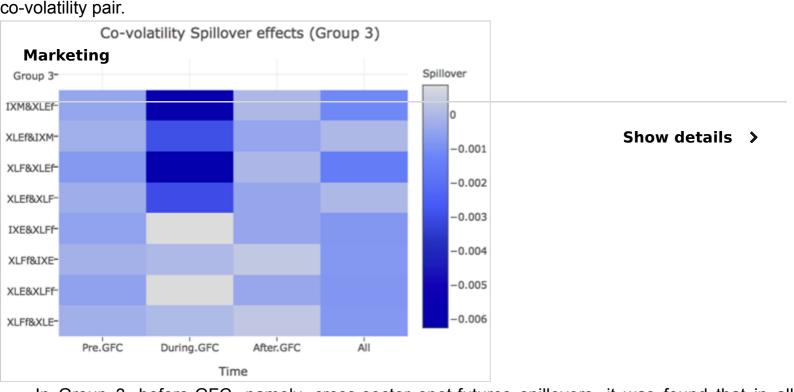
Pherences, during-GFC, it was found that in all cases, co-volatility spillovers we again statistically significant and negative. For each pair, the magnitude of the spillovers of the financial spot asset, namely IXM (Financial Select Sector Index) or XLF (financial ETF), on subsequent co-volatility between itself and its corresponding energy spot asset, namely IXE (Energy Select Sector Index) or XLE (energy ETF), was similar to the spillover effect of the energy spot asset on the same subset littles.

In Group 1, after-GFC, it was found that in all cases, co-volatility spillovers were statistically significant. For each pair, the spillovers of the financial spot asset, namely IXM (Financial Select Sector Index) or XLF (financial ETF), on subsequent co-volatility between itself **Showsdetails** ponding energy spot asset, namely IXE (Energy Select Sector Index) or XLE (energy ETF), was negative and greater than the positive spillovers of the energy spot asset on the same subsequent co-volatility pair.

In terms of the aggregation of the three periods for Group 1, it was found that in all cases, co-volatility spillovers were statistically significant and negative. For each pair, the magnitude of the spillovers of the financial spot asset, namely IXM (Financial Select Sector Index) or XLF (financial ETF), on subsequent co-volatility between itself and its corresponding energy spot asset, namely IXE (Energy Select Sector Index) or XLE (energy ETF), was less than the spillovers of the energy spot asset on the same subsequent co-volatility pair.



sub-periods, co-volatility spillovers were statistically significant. For the lone pair in this group, the magnitude of the spillovers of the financial futures asset, namely XLFf (financial ETF futures), on subsequent co-volatility between itself and its corresponding energy ETF futures, namely XLEf (energy ETF futures), was greater than the spillovers of the energy ETF futures on the came subsequent co-volatility pair. However, when the three sub-periods were combined, the opposite patterns were aled. In particular, the spillovers of the financial futures asset, namely XLFf mancial ETF futures), on subsequent co-volatility between itself and the energy ETF futures, namely XLEf (energy ETF futures), was less than the spillovers of the energy ETF futures on the same subsequent Statistics



In Group 3, before-GFC, namely, cross-sector spot-futures spillovers, it was found that in all cases, co-volatility spillovers were statistically significant and negative. For each pair, the magnitude of the spillover of the futures asset, namely XLFf (financial ETF futures) or XLEf (energy ETF futures), on subsequent co-volatility between itself and its corresponding cross-sector spot asset, namely IXE (Energy Select Sector Index) or XLE (energy ETF) and IXM (Financial Select Sector Index) or XLF Loading [MathJax]/jax/output/HTML-CSS/fonts/Gyre-Pagella/Size2/Regular/Main.js

(financial ETF), respectively, was greater than the spillovers of the spot asset on the same subsequent co-volatility pair.

In Group 3, during-GFC, it was found that co-volatility spillovers between XLEf (energy ETF futures) and XLF (financial ETF) or IXM (financial index), namely Cases 3.a.1 to 3.b.2, were

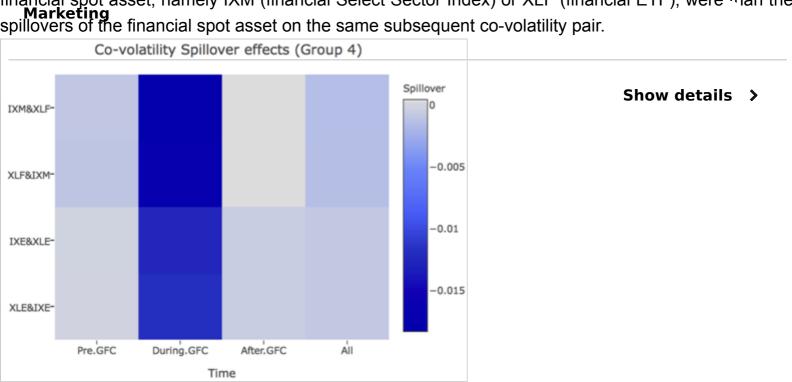
statistically significant and negative. For each pair, the magnitude of spillovers of XLEf (energy ETF futures) on subsequent co-volatility between itself and (https://www.cioopkiesetseotor/esp/wthatset, namely IXM (Financial Select Sector Index) or XLF (financial Financial Select Sector Index) or XLF (financial Financial Select Sector Index) or XLF (financial Financial Fin

In Group 3, after-GFC, it was found that in all cases, co-volatility spillovers were statistically

significant. For each pair, the magnitude of the spillover effect of XLFf (financial ETF futures) on

Subsequent co-volatility between itself and its corresponding cross-sector energy spot asset, namely IXE (Energy Select Sector Index) or XLE (energy ETF), was greater than the spillovers of the energy spot asset on the same subsequent co-volatility pair. However, the spillovers of XLEf (energy ETF futures) on subsequent co-volatility between itself and its corresponding cross-sector financial spot asset, namely IXM (financial Select Sector Index) or XLF (financial ETF), were positive and smaller than the negative spillovers of the financial spot asset on the same subsequent co-volatility pair. In Group 3, combining all three periods, it was found that in all cases, co-volatility spillovers were

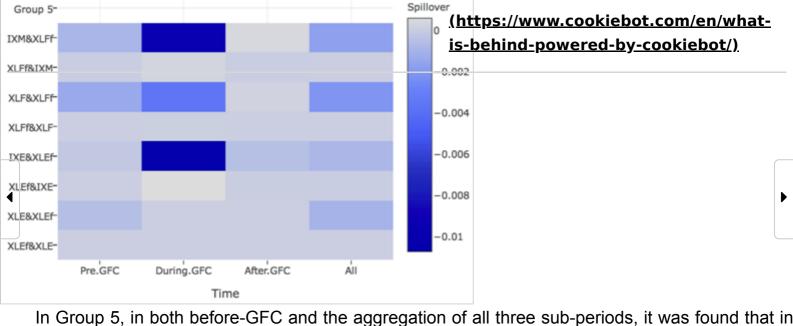
In Group 3, combining all three periods, it was found that in all cases, co-volatility spillovers were statistically significant and negative. For each pair, the magnitude of spillovers of XLFf (financial ETF futures) on Subsequent co-volatility between itself and its corresponding cross-sector engry spot asset, namely IXE (Energy Select Sector Index) or XLE (energy ETF), was similar to the spillovers of the season the same subsequent co-volatility pair. However, the spillovers of (energy ETF futures) on subsequent co-volatility between itself and its corresponding cross-sector financial spot asset, namely IXM (financial Select Sector Index) or XLF (financial ETF), were than the



In Group 4, it was found that in all cases, co-volatility spillovers were statistically significant over the four time periods. In terms of the magnitude of within-sector spot-spot co-volatility effects, the spillovers of IXM (Financial Select Sector Index) on subsequent co-volatility between itself and XLF

(financial ETF) were similar to the spillovers of XLF on the same subsequent co-volatility pair, namely Cases 4.a.1 and 4.a.2. This symmetry was also found for the pair involving co-volatility spillovers between the XLE (energy ETF) and IXE (energy index), namely Cases 4.b.1 and 4.b.2.

Co-volatility Spillover effects (Group 5)



all **Neses** sayvolatility spillovers were statistically significant. For each pair, the magnitude the spillovers of the futures asset, namely XLFf (financial ETF futures) and XLEf (energy ETF futures), on subsequent co-volatility between itself and its corresponding within-sector spot asset, namely IXM (Financial Select Sector Index) or XLF (financial ETF) and IXE (Energy Select Sector Index) or XLE (energy ETF), respectively, was greater than the spillovers of a given spot asset on the same subsequent co-volatility pair.

In Group 5, during-GFC and after-GFC, it was found that in all cases, co-volatility spillovers were statistically significant. In terms of the magnitude of within-sector spot-futures co-volatility e cts, for each pair, the spillovers of XLFf (financial ETF futures) on subsequent co-volatility between itself and its corresponding within-sector spot asset, namely IXM (Financial Select Sector Index) or XLF (financial ETF), were greater than the spillovers of the financial spot asset on the same subsequent co-volatility pair.

With regard to the within energy sector spot-futures co-volatility effect, the spillovers of XLEf (energy ETF futures) on subsequent co-volatility between itself and XLE (energy ETF) and the spillovers of XLE on the same subsequent co-volatility pair, namely Cases 5.d.1 and 5.d.2, were both significant, albeit, close to zero. However, the spillovers of XLEf (energy ETF futures) on subsequent co-volatility between itself and IXE (Energy index) were greater than the spillovers of the energy index on the same subsequent co-volatility pair, namely Cases 5.c.1 and 5.c.2.

All of the results pertaining to the five groups can be summarized by way of the six key findings given below. The terms symmetric and asymmetric, which are defined in terms of absolute values of spillover effects, are used for the first three findings. In particular, if a spillover pair is symmetric, it implies similar absolute values of spillover effects in both cases, based on casual empiricism. If a spillover effect pair is asymmetric, this indicates dissimilar absolute values of spillover pairs (in terms of casual empiricism in comparing the point estimates).

- Asymmetric spillover effects were found in all cases of spot-spot and futures-futures across sectors (see Groups 1 and 2). 2. Symmetric spillover effects were found in all cases of spot-spot between the financial ETF and financial index, as well as between the energy ETF and energy index in all periods (see Group 4). Asymmetric spillover effects were found in all cases of spot-futures ETF within sectors. Moreover, in
- all cases, spillover effects of ETF futures on its co-volatility typish/threwown costs in the co-volatility typish ty <u>is-behind-powered-by-cookiebot/)</u> than in the reverse case (see Group 5). 4. The co-volatility spillovers in all groups over all time periods are statistically significant, except for
- Cases 3.c.1 to 3.d.2 during-GFC. 5. Additionally, with the exception of the insignificant cases, the co-volatility spillovers are stronger during-GFC than for the other time periods (see Groups 1, 2 and 4). 6. In terms of the current relationship between the financial and energy sectors, the After-GFC
- spillovers are of greater relevance than the spillovers of the three sub-periods combined into a single sample.

6. Concluding Remarks

The primary purpose of the paper was to investigate the co-volatility spillovers within and across the**№.99.5117**gy and financial sectors in both their spot (namely, IXE, IXM, XLF and XLE) and f、∠res (namely, XLFf and XLEf) markets, by using "generated regressors" and a multivariate conditional volatility models namely diagonal BEKK. The daily data used in the empirical analysis are from 23 December 1998–22 April 2016. The dataset was analysed in its entirety and also subdivided into

financial spot asset, namely IXM (Financial Select Sector Index) or XLF (financial ETF), on sub**Magketing**p-volatility between itself and its corresponding energy spot asset, namely IXE Energy Select Sector Index) or XLE (energy ETF), was greater than the spillovers of the energy spot asset on the same subsequent co-volatility pair.

In Group 1, before and after the Global Financial Crisis, the magnitude of the spillove of the

three time periods, namely "before-GFC", "during-GFC" and "after-GFC".

However, during the GFC, the pattern changed dramatically. All of the spikers were stronger, and the spillovers of the financial spot asset on the subsequent co-volatility between itself and its corresponding energy spot asset was similar to the spillovers of the energy spot asset on the same subsequent co-volatility pair.

Other significant spillover patterns were also found between the financial ETF index and the energy ETF index in their spot-spot, spot-futures and futures-futures co-volatility, namely Groups 2 and 3, when combining all three periods. In terms of the within-sector spot-spot and spot-futures

markets, namely Groups 4 and 5, significant spillovers of ETF futures on subsequent co-volatility

between ETF and ETF futures were also found. It is well known that there is an intrinsic practical link between the financial and energy sectors, which can be analysed through their spillover effects, which are measures of how the shocks to

returns in different assets affect each other's subsequent volatility in both spot and futures markets. Moreover, it is apparent that there is an intrinsic relationship between the financial ETF and energy

ETF, both in their spot and futures markets. The empirical results showed that energy ETF and financial ETF have statistically significant co-volatility spillovers for all time periods. From a financial Loading [MathJax]/jax/output/HTML-CSS/fonts/Gyre-Pagella/Size2/Regular/Main.js managerial perspective, these empirical results suggest that financial and energy ETFs are suitable for constructing a financial portfolio from an optimal risk management perspective and also for dynamic hedging purposes. Failure to do so would miss out on optimal hedging and risk insurance for purposes of managing financial portfolios.

Author Contributions

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Chia-Lin Chang and Michael McAleer conceived the main ideas. Michael McAleer presented the technical analysis. Chien-Hsun Wang undertook the empirical analysis and wrote the first draft of the empirical results. Chia-Lin Chang and Michael McAleer edited the final draft.

Conflicts of Interest

The authors declare no conflict of interest.

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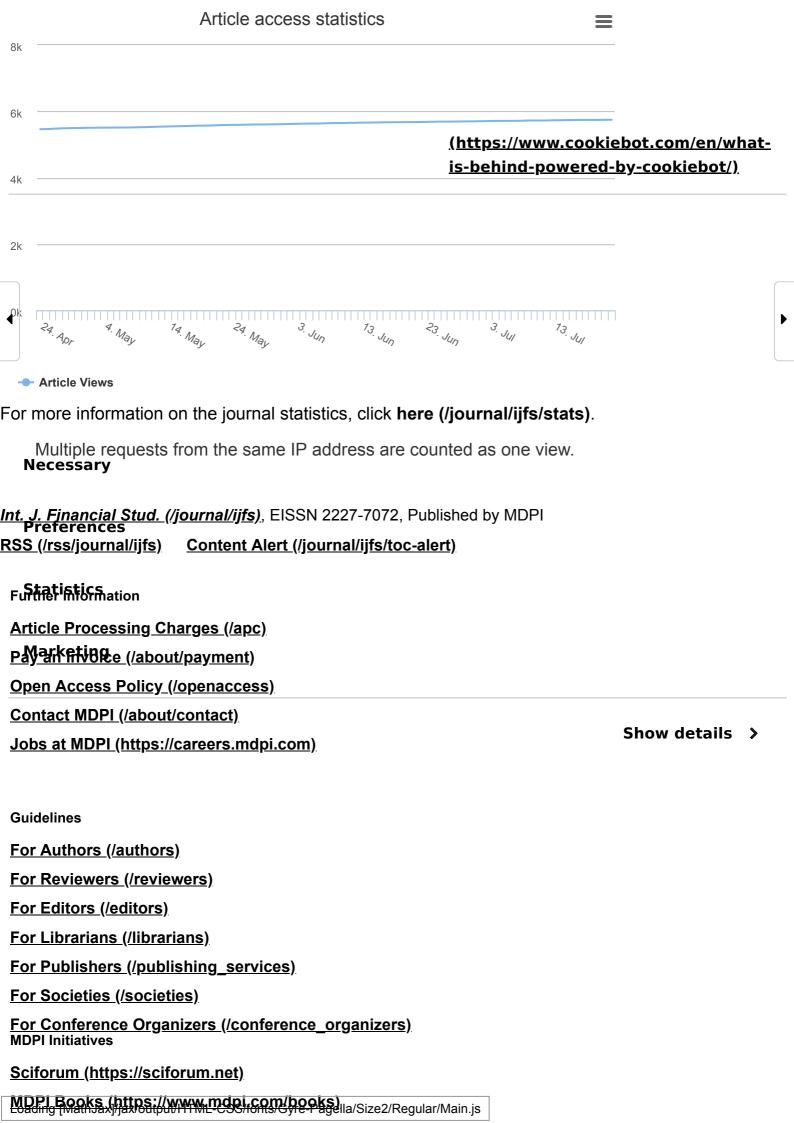
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