

Applied Economics >

Volume 40, 2008 - [Issue 16](#)

1,096 86

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Competition and efficiency in the Dutch life insurance industry

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Pages 2063-2084 | Published online: 11 Apr 2011

 Cite this article  <https://doi.org/10.1080/00036840600949298>

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Abstract

The lack of available prices in the Dutch life insurance industry makes competition an elusive concept that defies direct observation. Therefore, this article investigates competition by analysing several factors which may affect the competitive nature of a market and various indirect measurement approaches. After discussing various supply and demand factors which may constitute a so-called tight oligopoly, we establish the existence of scale economies and the importance of cost X-inefficiency, since severe competition would force firms to exploit available scale economies and to reduce X-inefficiencies. Both scale economies and X-inefficiencies turn out to be substantial, although more or less comparable to those found for insurers in other countries and to other financial institutions. Further, we apply the Boone indicator, a novel approach to measuring the effects of competition. This indicator points to limited competition in comparison to other sectors in the Netherlands. Further investigations of submarkets

should reveal where policy measures in order to promote competition might be appropriate.

†The views expressed in this article are personal and do not necessarily reflect those of CPB or DNB

Acknowledgements

The authors are grateful to Jan Boone, Marcel Canoy, Wil Dullemond, Aerd Houben, Marc Pomp, unknown referees and participants of the CPB Conference on ‘Competition in markets for life insurance’ (The Hague, April 27, 2005), the DNB Research Seminar (Amsterdam, 2005), the ‘Dynamics of Insurance Markets: Structure, Conduct, and Performance in the 21st Century’, Conference of the Wharton School/Journal of Banking and Finance (Philadelphia, USA, May 4–5, 2006) and Eurobanking 2006 (Dubrovnik, Croatia, May 21–24, 2006) for useful comments, Miriam Holman-Rijken and Leo Kranenburg for excellent research assistance and Marcel Eggenkamp and Usseb Karakhalil for providing the data.

Notes

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¹In terms of premiums as a percentage of GDP, the Dutch market is around 40% above the European weighted average.

²For an overview, see Bikker ([2004](#)) or Bikker and Bos ([2005](#)).

³For life insurances, a second motive is the accumulation of assets. Some countries see many buyers of annuities eventually cashing out their contracts rather than annuitizing.

⁴In the Netherlands, health insurance is part of nonlife insurance, whereas in Anglo-Saxon countries, health insurance is seen as part of life insurance.

⁵A typical endowment insurance policy pays a given amount at a given date if a given person is still alive, or earlier when he or she passes away. Of course, there are many variants to these archetypes.

⁶The fiscal regime change might cause a structural break. However, re-estimation of our model for two sub-periods, before and after the change, did not give different results.

⁷For a fuller discussion we refer to CPB ([2005](#)). See also Kamerschen ([2004](#)).

⁸For a detailed analysis of the various effects we refer to CPB ([2003](#)).

⁹Concentration ratios are discussed in Bikker and Haaf ([2002](#)). where s_i represents the market share of firm i .

¹⁰In 1996, Japanese entrance increased sharply due to a structural change.

¹¹Acquisition costs are marketing costs and sales costs, which include commissions to insurance agents.

¹²Incidentally, a new Dutch Financial Services Act (Wet Financiële Dienstverlening) has come into force at the begin of 2006, pressing for more transparency in this market, which may also work to improve competition in this submarket.

¹³See Consumentenbond, 2004, Consumentengeldgids (Personal finance guide), September, 34-37.

¹⁴This interpretation would be different in a market with only few firms, so that further consolidation would be impossible. Further, this interpretation would also change when new entrees incur unfavourable scale effects during the initial phase of their growth path.

¹⁵Note that sometimes scale economies are defined by the reciprocal of Equation [3](#), see, for instance, Baumol et al . ([1982](#), p. 21) and Resti ([1997](#)).

¹⁶The first stochastic frontier function for production was independently proposed by Aigner et al . ([1977](#)) and Meeusen and van den Broeck ([1977](#)). Schmidt and Lovell ([1979](#)) presented its dual as a stochastic cost frontier function.

¹⁷This expression relies upon the predicted value of the unobservable, u_{it} , which can be calculated from expectations of u_{it} , conditional upon the observed values of v_{it} and

u_{it} , (Battese and Coelli, [1992](#), 1993, 1995).

¹⁸Note that the $E(c_{it} | u_{it} , X)$ differs from actual costs, c_{it} , due to v_{it} .

¹⁹An alternative definition would be the inverse of EFF_{it} , $INEFF_{it} = \exp(u_{it})$, which is bounded between 1 and ∞ .

²⁰See Boone and Weigand in CPB ([2000](#)) and Boone ([2001](#), 2004).

²¹More competition can force firms to consolidate (see our scale economies discussion). Claessens and Laeven ([2004](#)) found in a world wide study on banking that concentration was positively instead of negatively related to competition.

²²Suppose that the negative profit firms are price fighters. In a well-functioning market the price fighters will-influence profitability of the other firms.

²³Some insurance firms can approximate their value added by comparing their embedded value over time. These data are not publicly available.

²⁴The definition of production of life insurance firms is discussed further in subsection 'cost X-inefficiency'.

²⁵The price of management, or wages, has been excluded by applying the two standard properties of cost functions, namely linear homogeneity in the input prices and cost exhaustion (Jorgenson, [1986](#)).

²⁶Of course, the accuracy of this optimal size is limited, as its calculated location lies far out of our sample range.

²⁷This figure is based on the OLS estimates, which provides the starting values of the numerical optimization procedure. As OLS minimizes the errors terms and maximizes the degree of fit, the latter will be lower in the SCF model.

²⁸This measure can be defined as $\frac{p_i}{mc_i}$ where p_i denotes the firm's equilibrium output price and mc_i its marginal cost.

²⁹ISIS data concern both domestic and foreign activities. Pure domestic figures would be more precise but are not available.

³⁰For instance, firms in the Netherlands use more agents as selling channel than those in other countries (CEDA, [2004](#), p. 144).

³¹A similar picture emerges from figures of CEDA ([2004](#)), p. 198.

³²This lagging adjustment of profitability does not disturb the international comparison, as this limitation holds also for the foreign data.

³³Note that the variable cost may change over the size classes due to scale efficiency (just as the marginal cost may do), so that the average variable cost may differ from the marginal cost. Apart from this theoretical dissimilarity, these variables are also measured differently in practice.

³⁴We have also estimated random effect models for profits ([Table 8](#)) and markets shares ([Table 9](#)). Their coefficients were quite similar to those of the fixed effect models, with even slightly higher values and higher levels of significance. This suggests that the estimates presented in [Tables 8](#) and [9](#) are quite robust. We tested for random effect using the Hausman test, but this test appeared to be undefined, suffering from the 'small sample problem'. All models include year dummies, also not shown in the tables.

³⁵The value of the Boone-indicator in these estimations is around -0.85 . Results can be obtained from the authors.

³⁶The elasticity of this variable is the coefficient (0.45) times the average of the unit-linked fund ratio (0.33; see [Table 3](#)), so 0.15.

³⁷In the basic model, the β_1 values for mc are lower than for average variable costs (namely around -1) and for one year even not significant, see [Table A2](#) in Appendix 2.

³⁸The elasticity, the first derivative of the auxiliary equation in logs, is $-0.37 + 0.01 \times 2 \times$ average production in logarithms. For the auxiliary model in natural values it is equal to $mc \text{ production} \times (\text{average production}/\text{average mc}) = (-0.134e^{-7} + (0.249e^{-14} \times 247707.4 \times 2) \times 247707/0.18$.

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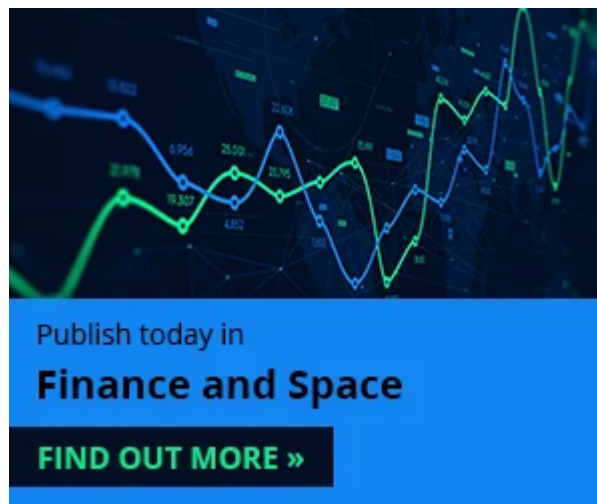
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