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Letter

The need for better analysis of observational studies in orthopedics

A retrospective study of elbow fractures in children

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Abstract

Background and purpose The conventional statistical methods employed in observational studies in orthopedics require the fundamental assumption that the outcomes are independent. However, fractures treated by the same surgeon cannot be regarded as being independent of each other and should be nested in the statistical analysis. If the effect on outcome of early rather than delayed surgery depends on the severity of the fracture, the results of a level model

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Results We found that 25% of the variance in outcome could be attributed to between-surgeon variance. We identified an interaction between the surgeons' experience and the severity of the fractures that influenced the conclusions. The variable “number of pins” was not significant in the 2-level model ($p = 0.07$), while the ordinary least-squares analysis gave a result that was statistically significant ($p = 0.01$).

Interpretation Researchers should consider the need for a 2-level model and the presence of interactions. Standard statistical methods might lead to wrong conclusions.

Supracondylar humerus fractures are the most common elbow injuries in children who require surgery. Percutaneous pinning has become the method of choice in most clinics, and severe complications are rare (Otsuka and Kasser [1997](#)). Recent reports have suggested that delay of surgery until the next day is safe (Iyengar et al. [1999](#), Mehlman et al. [2001](#), Leet et al. [2002](#), Gupta et al. [2004](#)). However, vascular injuries and compartment syndromes still occur (Ramachandran et al. [2008](#)), and some authors recommend treating these fractures as early as possible (Walmsley et al. [2006](#)). The statistical methods employed in these publications are t-test, the chi-square test, and simple and multiple regression analyses. These methods require the fundamental assumption that the observed outcomes are independent, which implies that none of the fractures should have been operated by the same surgeon, and the distribution of fractures among surgeons should be random. This is rare, if not impossible, in observational studies on orthopedic surgery.

In the social sciences, there has been a gradual awareness of the inability of conventional statistical methods to analyze the complexity of human interaction. Researchers in public health have repeatedly encouraged the integration of social science

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approach (also called hierarchical modeling) accounts for potential correlations by modeling intercepts and regression coefficients as random. The intraclass correlation coefficient (ICC) expresses the amount of dependency among observations and is calculated to decide whether a multilevel analysis is appropriate. The ICC can take values from 0 to 1. A non-zero value of ICC implies that the observations are not uncorrelated and that there is a need for multilevel modeling.

Furthermore, orthopedic researchers should consider the existence of interaction, also known as effect modification (Moyé [2006](#)). An interaction is defined as a factor that modifies the independent factor under study. This is analytically more complex than simple confounding. A confounder has the same effect on outcome for all values of the other independent variables studied. Interactions reflect that the effect of one variable depends on the values of one or more other variables. For example, the influence of the surgeon's experience on outcome could be stronger for severe fractures than for less complicated ones. In such cases, a statistical model with interactions should be tested.

Here we describe to the orthopedic community the concept of multilevel modeling and interactions as necessary statistical tools in observational studies. We show that the conventional statistical methods currently employed in retrospective reports may yield misleading results.

Patients and methods

The study protocol was approved by the Regional Ethics Committee on October 25, 2007 (registration number 1.2007.2093). The patient population consisted of all children who underwent reduction and pinning of a displaced supracondylar humerus fracture in our institution between 1999 and 2006. The patients were identified in our



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Table 1. Demographic and fracture characteristics of 112 supracondylar humerus fractures in 112 children



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Radiographic analysis

Standard AP and lateral radiographs were used. The Gartland classification of the lateral film was recorded by PHR (Gartland [1959](#)). This classification has good intra- and interobserver reliability (Barton et al. [2001](#)).

Clinical examination

All patients were given written information about the study and were invited to attend a follow-up clinic. 3 patients had moved abroad or to a different part of the country. 9 patients were lost to follow-up, and 22 patients refused to participate, leaving 78 patients available for clinical examination on average 4.3 (1.5–9) years after admission. The surgeon's experience was not revealed to the investigators examining the patients (PHR, EAS, and IS). The cubitus angle and range of motion were measured clinically using a goniometer. The Flynn criteria (Flynn et al. [1974](#)) were used to assess the functional and cosmetic outcome. Neurovascular status was assessed clinically and the patients completed the Quick-DASH score (www.dash.iwh.on.ca [2009](#)) with the help of their parents. The overall satisfaction of the end-result was assessed using a VAS score where zero was the worst imaginable result and 10 was the best (equivalent to the uninjured arm).

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A hierarchical linear modeling approach with 2 levels (patient and surgeon) was chosen to model the main outcome variable VAS score. Even though the VAS score was a highly skewed variable, a 2-level linear model (Fitzmaurice et al. [2004](#)) was estimated. Simple analyses were then used to assess the association between the outcome variable and the main predictors. 4 variables were considered to be main predictors for the VAS score. The Gartland classification and the surgeon's level of experience (either consultant or resident) were considered to be predictor variables based on the notions that more displaced fractures have more complications and that more experienced surgeons have better results. The variable night (between midnight and 8 a.m.) or day was selected based on the tendency to perform fewer operations during the night to reduce the risk of adverse outcome (Rothschild et al. [2009](#)). There is no consensus on the number of pins needed for adequate fixation. Some authors recommend the use of 3 pins in unstable fractures (Gordon et al. [2001](#), Vlahovic and Bumci [2002](#)). We therefore included the number of pins as a predictor variable in the model. Time to surgery, preoperative neurovascular status, age, and sex were considered as confounders.

Technically, to account for interactions a new variable, which is the product of 2 (usually) independent variables, is introduced into the model. One should not try out all possible interactions since this increases the risk of spurious results. Thus, when building the multiple model, we only considered variables that we thought might produce an interaction. We decided that the surgeon's level of experience (consultant or resident) and time to surgery might have different effects on outcome depending on the severity of the fracture. We therefore included these 2 interaction terms in the original multiple model. The R^2 , quantifying the proportion of outcome variation explained, was used for model reduction. Akaike's information criterion was applied for covariance model selection (Akaike [1974](#)). Even though the hypothesis about normally distributed residuals could not be accepted, the qq plot for residuals demonstrated an almost s

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fractures were operated earlier and more often by consultants, and had more complications than Gartland 2 fractures.

Table 2. Postoperative complications recorded for the 112 fractures



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11 of the 78 elbows available for clinical examination had a cubitus varus malunion. The mean VAS score was lower in patients who developed cubitus varus (7.7 vs. 9.6) than in patients who had a positive carrying angle ($p < 0.001$, Mann-Whitney test). 20 patients had a positive Quick-DASH score (Table 3).

Table 3. Clinical outcome of 78 fractures at follow-up, mean 4.3 years after surgery



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The overall amount of variation in outcome explained by our model was 18%. By estimating the unconditional model, the ICC was calculated to be 0.25; i.e. 25% of the total variance in VAS score could be accounted for by the surgeon level. This suggests that an ordinary least-squares (OLS) regression analysis would probably yield misleading results. The simple analyses did not show any statistically significant relationships between the 4 predictors and VAS score. However, 10–26% of the explainable variation in outcome was explained by the predictors, which was sufficiently high to be able to include all considered predictors in the multiple model (Table 4).

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In the OLS model without interactions, there was a positive correlation between number of pins and patient satisfaction ($p = 0.02$), while in the 2-level model with interaction terms included this correlation was not significant ($p = 0.07$). Neither the surgeon's experience nor the severity of the fracture influenced the VAS score significantly. The combination (interaction) of the two, with marginally significant effect on the outcome ($p = 0.05$), influenced the conclusions sufficiently to be included in the final model.

Although the 2-level model corrected the problem of underestimated standard errors in OLS regression, the results did not differ much in the 2 models. Only the number of pins was no longer statistically significant in the 2-level model. If we had used only a simple OLS model, we would have reported number of pins as an explanatory variable.

Discussion

Our point is not that the last word about treatment of supracondylar fractures has been said, but that simple statistical models do not suffice. We found that the clinical outcome after pinning of displaced supracondylar humerus fractures in children is good or excellent in most cases, which is consistent with previous reports (Foead et al. [2004](#), Devkota et al. [2008](#)). The few complaints were mainly cosmetic, without loss of function. This is a difficult starting point when searching for factors to improve treatment, and we could only explain 18% of the variance in outcome. Even so, the use of statistical analyses that adhere more closely to reality than a simple OLS model changed conclusions. We have considered the most commonly used explanatory variables for inferior outcome and found that an interaction variable affected

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without interactions has been used—and where the conclusions have had an impact on clinical practice—should be reviewed using these statistical methods.

PHR initiated and designed the study, collected data, organized the follow-up clinic, examined patients, and wrote the manuscript. EAS and IS examined patients and reviewed the manuscript. JSB built the statistical models and revised the manuscript. PG contributed to the study design and to critical revision of manuscript.

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No competing interests declared.

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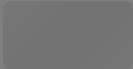

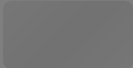
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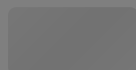
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